Module 7: Mixed autoregressive moving average (ARMA) practice problems
(The attached PDF file has better formatting.)
** Exercise 7.1: Mixed autoregressive moving average (ARMA) process
An $\operatorname{ARMA}(1,1)$ process has $\sigma^{2}=1, \theta=-0.4$, and $\phi=0.8$.
A. What is the value of $\gamma_{0}$ ?
B. What is the value of $\gamma_{1}$ ?
C. What is the value of $\rho_{1}$ ?
D. What is the value of $\rho_{2}$ ?

ARMA $(1,1)$ means the process has $1 \phi$ parameter and $1 \theta$ parameter. Know equations 4.4.3, 4.4.4, and 4.4.5 on page 78. These seem complex at first. Review the derivation in the textbook, so you recall the formulas.

After period 1, $\gamma_{k}$ and $\rho_{\mathrm{k}}$ have exponential decay. $\rho_{1}$ is $\gamma_{1} / \gamma_{0 .} \gamma_{1}$ is a simple function of $\gamma_{0} \phi_{1}$, and $\theta$. The final exam problems test $\rho_{1}$ most frequently.

Part A: For an ARMA(1,1) process, $\gamma_{0}=\frac{\left(1-2 \phi \theta+\theta^{2}\right)}{1-\phi^{2}} \sigma_{e}^{2}$
$=\left(1-2 \times 0.8 \times-0.4+(-0.4)^{2}\right) /\left(1-(0.8)^{2}\right)=5$
(See Cryer and Chan, chapter 4, page 78, equation 4.4.4)
Part B: $\gamma_{1}=\phi \gamma_{0}-\theta \sigma_{\varepsilon}^{2}=0.8 \times 5-(-0.4) \times 1=4.4$
See Cryer and Chan, chapter 4, page 78, equation 4.4.3
Jacob: How do we derive these formulas?
Rachel: $\gamma_{0}$ is the variance of $Y_{t}=E\left(Y_{t}^{2}\right)$.
To avoid superfluous parameters, subtract the mean from all values of the time series.
The variance of $Y_{t}$, the covariances, and the autocorrelations do not change, and we don't have to include a $\mu$ parameter in the derivations.

For an $\operatorname{ARMA}(1,1)$ process with a mean of zero, $\mathrm{Y}_{\mathrm{t}}=\phi \times \mathrm{Y}_{\mathrm{t}-1}+\epsilon_{\mathrm{t}}-\theta \times \epsilon_{\mathrm{t}-1} \Rightarrow$
$\gamma_{0}=E\left(Y_{t}^{2}\right)=E\left(Y_{t} \times\left(\phi \times Y_{t-1}+\epsilon_{t}-\theta \times \epsilon_{t-1}\right)=\phi \times E\left(Y_{t} \times Y_{t-1}\right)+E\left(\epsilon_{t} \times Y_{t}\right)-\theta \times E\left(\epsilon_{t-1} \times Y_{t}\right) \Rightarrow\right.$

$$
\gamma_{0}=\left(\phi \times \gamma_{1}\right)+E\left(\epsilon_{t} \times Y_{t}\right)-\theta \times E\left(\epsilon_{t-1} \times Y_{t}\right)
$$

Also: $\gamma_{1}=E\left(Y_{t-1} \times Y_{t}\right)=E\left(Y_{t-1} \times\left(\phi \times Y_{t-1}+\epsilon_{t}-\theta \times \epsilon_{t-1}\right)=\phi \times E\left(Y_{t-1} \times Y_{t-1}\right)+E\left(\epsilon_{t} \times Y_{t-1}\right)-\theta \times E\left(\epsilon_{t-1} \times Y_{t-1}\right)\right.$.
For a stationary time series, $E\left(Y_{t-1} \times Y_{t-1}\right)=E\left(Y_{t} \times Y_{t}\right)=\gamma_{0}$ and $E\left(\epsilon_{t-1} \times Y_{t-1}\right)=E\left(\epsilon_{t} \times Y_{t}\right)$, so

$$
\gamma_{1}=\left(\phi \times \gamma_{0}\right)+E\left(\epsilon_{t} \times Y_{t-1}\right)-\theta \times E\left(\epsilon_{t} \times Y_{t}\right)
$$

If we can solve for $E\left(\epsilon_{t} \times Y_{t}\right)$ and $E\left(\epsilon_{t-1} \times Y_{t}\right)$ in terms of the $\operatorname{ARMA}(1,1)$ parameters $\phi, \theta$, and $\sigma^{2}{ }_{t}$, we have two equations in two unknowns ( $\gamma_{0}$ and $\gamma_{1}$ ).
$\epsilon_{\mathrm{t}}$ is uncorrelated with $\mathrm{Y}_{\mathrm{t}-1}$ and with $\epsilon_{\mathrm{t}-1}$, so

$$
\mathrm{E}\left(\epsilon_{\mathrm{t}} \times \mathrm{Y}_{\mathrm{t}}\right)=\mathrm{E}\left(\epsilon_{\mathrm{t}} \times\left(\phi \times \mathrm{Y}_{\mathrm{t}-1}+\epsilon_{\mathrm{t}}-\theta \times \epsilon_{\mathrm{t}-1}\right)\right)=0+\sigma_{\mathrm{t}}^{2}+0=\sigma_{\mathrm{t}}^{2}
$$

Similarly, $\epsilon_{\mathrm{t}-1}$ is uncorrelated with $\epsilon_{\mathrm{t}}$ and the correlation of $\epsilon_{\mathrm{t}-1}$ with $\mathrm{Y}_{\mathrm{t}-1}$ is $\sigma^{2}{ }_{\mathrm{t}}$, so

$$
E\left(\epsilon_{t-1} \times Y_{t}\right)=E\left(\epsilon_{t-1} \times\left(\phi \times Y_{t-1}+\epsilon_{t}-\theta \times \epsilon_{t-1}\right)\right)=\phi \times \sigma_{t}^{2}+0-\theta \times \sigma_{t}^{2}=(\phi-\theta) \times \sigma_{t}^{2}
$$

We now have two equations in two unknowns:

$$
\begin{gathered}
\gamma_{0}=\left(\phi \times \gamma_{1}\right)+\sigma_{t}^{2}-\left(\theta \times(\phi-\theta) \times \sigma_{t}^{2}\right) \Rightarrow \\
\gamma_{0}=\left(\phi \times \gamma_{1}\right)+\left([1-\theta \times(\phi-\theta)] \times \sigma_{t}^{2}\right) \\
\gamma_{1}=\left(\phi \times \gamma_{0}\right)-\left(\theta \times \sigma_{t}^{2}\right)
\end{gathered}
$$

We use the expression for $\gamma_{1}$ in the formula for $\gamma_{0}$ to get

$$
\begin{gathered}
\gamma_{0}=\left(\phi \times\left[\left(\phi \times \gamma_{0}\right)-\left(\theta \times \sigma_{\mathrm{t}}^{2}\right)\right]\right)+\left([1-\theta \times(\phi-\theta)] \times \sigma_{\mathrm{t}}^{2}\right) \Rightarrow \\
\left.\left(1-\phi^{2}\right) \times \gamma_{0}=\left(\phi \times\left(\theta \times \sigma_{\mathrm{t}}^{2}\right)\right]\right)+\left([1-\theta \times(\phi-\theta)] \times \sigma_{\mathrm{t}}^{2}\right) \Rightarrow \\
\left.\gamma_{0}=\left(1-2 \theta \phi-\theta^{2}\right) \times \sigma_{\mathrm{t}}^{2}\right) /\left(1-\phi^{2}\right)
\end{gathered}
$$

We express $\gamma_{1}$ as a function of $\phi, \theta$, and $\sigma^{2}$ :

$$
\begin{gathered}
\left.\gamma_{1}=\left(\phi \times \gamma_{0}\right)-\left(\theta \times \sigma_{\mathrm{t}}^{2}\right)=\left\{\left[\phi \times\left(1-2 \theta \phi-\theta^{2}\right)\right]-\left[\theta \times\left(1-\phi^{2}\right)\right]\right\} \times \sigma^{2}{ }_{\mathrm{t}}\right) /\left(1-\phi^{2}\right) \Rightarrow \\
\gamma_{1}=\left[\left(\phi-\theta-2 \theta \phi^{2}+\theta \phi^{2}\right) \times\right]{\sigma_{\mathrm{t}}}^{2} /\left(1-\phi^{2}\right) \Rightarrow \\
\gamma_{1}=\left[(\phi-\theta) \times\left(1-2 \theta \phi^{2}+\theta \phi^{2}\right)\right] \times \sigma_{\mathrm{t}}^{2} /\left(1-\phi^{2}\right)
\end{gathered}
$$

We express $\rho_{1}$ as a function of $\phi$ and $\theta$ :
$\rho_{1}=\gamma_{1} / \gamma_{0}$. Both $\gamma_{0}$ and $\gamma_{1}$ have the multiplicative term $\sigma^{2}{ }_{t} /\left(1-\phi^{2}\right)$, which cancels out of the ratio for $\rho_{1}$ :

$$
\rho_{1}=\left[(\phi-\theta) \times\left(1-2 \theta \phi^{2}+\theta \phi^{2}\right)\right] /\left(1-2 \theta \phi-\theta^{2}\right)
$$

The moving average effect dies out after one period, and the autoregressive effect decays exponentially, so

$$
\begin{array}{r}
\rho_{\mathrm{k}}=\rho_{\mathrm{k}-1} \times \phi \Rightarrow \\
\rho_{\mathrm{k}}=\rho_{1} \times \phi^{\mathrm{k}-1}
\end{array}
$$

See Cryer and Chan, chapter 4, pages 77-78, equations 4.4.2-4.4.4. Cryer and Chan write the formulas for $\gamma_{0}, \gamma_{1}$, and $\rho_{k}$. Rachel shows the derivation in more detail. Once you have read Rachel's dialogue, you can see the intuition for the expressions. The equations appears on final exam problems, so follow Rachel's derivation.

Part C: For an $\operatorname{ARMA}(1,1)$ process: $\rho_{k}=\frac{(1-\phi \theta)(\phi-\theta)}{1-2 \phi \theta-\theta^{2}} \phi^{k-1}$

Since we have solved for $\gamma_{0}$ and $\gamma_{1}$, we use the relation $\rho_{1}=\gamma_{1} / \gamma_{0}=4.4 / 5=0.88$
In practice, you use the autocorrelations $\rho_{\mathrm{k}}$ more frequently than the autocovariances $\gamma_{\mathrm{k}}$.
Part D: The autocorrelations $\rho_{\mathrm{k}}$ decline geometrically at a rate $\phi \Rightarrow \rho_{2}=\rho_{1} \times \phi=0.88 \times 0.8=0.704$
In the module on method of moments, we start with sample autocorrelations of 0.88 and 0.704 and derive estimates of the $\phi$ and $\theta$ parameters.
** Exercise 7.2: Moving average and autoregressive processes
An autoregressive process with mean zero is $Y_{t}=\phi Y_{t-1}+\epsilon_{t}$
Re-write this time series as a moving average process of infinite lag.

## Solution 7.2:

$$
\begin{aligned}
Y_{t} & =e_{t}+\phi Y_{t-1} \\
& =e_{t}+\phi\left(e_{t-1}+\phi Y_{t-2}\right) \\
& =e_{t}+\phi e_{t-1}+\phi^{2}\left(e_{t-2}+\phi Y_{t-3}\right)
\end{aligned}
$$

Continue expanding to eliminate all the $Y_{t}$ and remain with an infinite series of $\epsilon$ 's.
$Y_{t}=e_{t}+\phi e_{t-1}+\phi^{2} e_{t-2}+\phi^{3} \epsilon_{3}+\ldots$
See Cryer and Chan, chapter 4, page 70, equation 4.3.8
Cryer and Chan refer to the coefficients of the $\epsilon$ 's as $\psi$ weights or filter representation.
Jacob: Why would we want to express an AR(1) process as a moving average process of infinite lag?
Rachel: The values of an $\operatorname{AR}(1)$ process are autocorrelated. The $\epsilon$ terms are independent, so we can easily compute the variance of the autoregressive process as

$$
\text { variance }\left(Y_{t}\right)=\left(1+\phi+\phi^{2}+\phi^{3}+\ldots\right) \times \sigma_{t}^{2}=\sigma_{t}^{2} /\left(1-\phi^{2}\right) .
$$

An ARMA $(1,1)$ process has $\rho_{1}=0$. Which of the following is true?
A. $\quad \phi_{1}=\theta_{1}$ or $\phi_{1} \times \theta_{1}=1$
B. $\phi_{1}=\theta_{1}$ and $\phi_{1} \times \theta_{1}=1$
C. $\phi_{1}=-\theta_{1}$ or $\phi_{1} \times \theta_{1}=-1$
D. $\phi_{1}=-\theta_{1}$ and $\phi_{1} \times \theta_{1}=-1$
E. $\phi_{1}=\theta_{1}$ and $\phi_{1} \times \theta_{1}=-1$

Answer 7.3: A
For an $\operatorname{ARMA}(1,1)$ process:
$\rho_{k}=\frac{(1-\theta \phi)(\phi-\theta)}{1-2 \theta \phi+\theta^{2}} \phi^{k-1}$
for $k \geq 1$

Notation: For $\operatorname{AR}(1), \mathrm{MA}(1)$, and $\operatorname{ARMA}(1,1)$ processes, Cryer and Chan drop the subscripts on the parameters, using $\phi$ and $\theta$ instead of $\phi_{1}$ and $\theta_{1}$. The final exam problems may use either notation.

Intuition: If the residual in period t increases one unit:

- Moving average: $\theta_{1}$ causes the forecast to decrease $\theta_{1}$ units
- Autoregressive: $\phi_{1}$ causes the forecast to increase $\phi_{1}$ units

If $\phi_{1}=\theta_{1}$, these two effects offset each other. A change in the period $t$ value does not affect the period $t+1$ value, so $\rho_{1}=0$.
$\theta_{1}$ and $1 / \theta_{1}$ produce the same autocorrelation function.
$\phi_{1}=\theta_{1}$ has the same effect as $\phi_{1}=1 / \theta_{1}$ or $\phi_{1} \times \theta_{1}=1$.

