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### **1. Introduction**

The report excerpt from my personal master degree's paper in 2010.

In this report, I assume that there are two potential assets in policyholders' portfolio; one with high risk and high return and the other one with low risk and low return. The utility function of the policyholder is assumed to follow a poIr utility. I consider the asset allocation effect on the guaranteed cost for a VA with guaranteed minimum withdrawal benefits, finding that the guaranteed cost will increase 12% compared with a specific underling asset.

The model effect of the asset return process is also examined by considering two different asset processes, the lognormal model and ARIMA-GARCH model. The solution of dynamic programming problem is solved by the numerical approach proposed by Huang (2009). Finally I get the conclusion which the guaranteed cost given by the ARIMA-GARCH model is greater than the lognormal model.

# 2 .Methodology

In this chapter, I introduce and explain our design of guaranteed minimum benefit, utility function model of the policyholder and the optimal strategy.

# 2.1 Design of Guaranteed Minimum

# Benefit

I consider a T-year variable annuity with guaranteed minimum benefit. For simplicity, I assume that there are two assets, and the policyholder can make asset allocation decision on his portfolio in the beginning of each year. One of two assets has high return and high risk and the other has low return and low risk. The following are the definitions of the symbols:

 $S_{1,t}$ : the value of high return and high risk asset at time t  $S_{2,t}$ : the value of low return and low risk asset at time t  $w_t$ : the light invested in high return and high risk asset in the beginning of year t  $F_t$ : the market value of separate account in the end of year t  $G_t$ : the guaranteed value of separate account at the end of year t  $P_t$ : the premium at time t m: the fee ratio of account d: withdrawal rate, a percentage of single premium

#### **Guaranteed minimum maturity Benefit**

I consider a T-year variable annuity with guaranteed minimum maturity benefit. After T year, the policyholder will get the account value or the guaranteed value whichever is greater.

The market value of account at time t + 1 is

$$F_{t+1} = (F_t + P_t)(1 - m)(w_t \frac{S_{1,t+1}}{S_{1,t}} + (1 - w_t)\frac{S_{2,t+1}}{S_{2,t}})$$

At the end of year T, the policyholder can receive:

 $\max \{G_T, F_T\}$ 

#### Guaranteed minimum withdrawal Benefit

I consider a T-year variable annuity with single premium guaranteed minimum withdrawal benefit. The policyholders will withdraw a percentage of the single premium at the end of each year until the guaranteed value becomes zero.

The market value of account at time t + 1:

$$F_{t+1} = (F_t - d * P_0)(w_t \frac{S_{1,t+1}}{S_{1,t}} + (1 - w_t) \frac{S_{2,t+1}}{S_{2,t}})$$

The guaranteed value at time t:

$$\begin{cases} G_t = G_{t-1} - d * P_0 & t = 1,2,3,4, \dots T \\ G_0 = P_0 & t = 0 \end{cases}$$

At the end of year T, the policyholder can receive:

 $\max \{G_T, F_T\}$ 

### **2.2 Asset Model**

In order to avoid asset model error, I assume that assets are subject to two different models, log-normal model and the Autoregressive Integrated Moving Average-Generalized Autoregressive Conditional Heteroscedastic (ARIMA-GARCH) model.

#### Lognormal model

Lognormal model assumes that return of asset after taking logarithmic will follow Normal distribution. The dynamic processes of two assets are

 $dlog (S_{t,k}) = \mu_k log (S_{t,1}) d_{t+} \sigma_k log (S_{t,k}) dB_{t,k}$ 

k=1,2 '1' represent high risk asset and '2' represent low risk asset  $S_{t,k}$ : asset value at time t  $\mu_k$ : return of asset at P measure  $dB_{t,k}$ :standard Brownian motion  $dB_{t,k} \sim N(0,t)$  $\sigma_k$ : the volitiity of the asset after the logarithmic

#### Autoregressive Integrated Moving Average-Generalized Autoregressive

#### Conditional Heteroscedastic (ARIMA-GARCH) model

First, Morgan (1976) shows that the return of equity has a heterogeneous phenomenon which represents the volatile of the equity return will change with time. Mandelbrot (1963) and Fama (1965) propose the distribution of financial time series has the characteristics of leptokurtic, fat tail and non-lognormal distribution; and its volatile has the phenomenon of volatility clustering. As a result, Engle (1982) develops ARCH (Autoregression Conditional Heteroskedaticity) model. The main characteristic of this model is taking the residual of lag periods as a conditional variance; the model later gets the empirical support from the market data in UK. After this, Bollerslev (1986) expands ARCH model to GARCH model (Generalized Autoregression Conditional Heteroskedaticity), and shows GARCH model can capture all the features of the volatile. Therefore I use this model to fit the linked assets in our study. The asset process of ARIMA(p, e, q)-GARCH(j, k) model :

$$\Delta^{d} y_{t,k} = a_{0,k} + \sum_{i=1}^{p} \phi_{i,k} \Delta^{d} y_{t-i,k} + \sum_{i=1}^{q} \theta_{i} \Delta^{d} \varepsilon_{t-i,k} + \varepsilon_{t,k}$$

 $\epsilon_{t,k} \sim N(0, \sigma_{t,k})$ 

$${\sigma_{t,k}}^2 = \omega_{0,k} + \sum_{i=1}^{j} \rho_{i\sigma_{t-i,k}{}^2} + \sum_{i=1}^{k} \gamma_i \epsilon_{t-i,k}{}^2$$

k='1','2' '1' represents high risk asset and '2' represents low risk asset  $S_{i,j}$ 

$$y_t: \log\left(\frac{S_{t,k}}{S_{(t-1),k}}\right)$$

 $S_{t,k}$ : the asset value at time t

p: the order of the autoregressive part

e: the order of the integrated part

q: the order of the moving average part

 $\epsilon_{t,k}$  : White noise of asset k and follow normal distribution

 $\sigma_{t,k}$  the volitility of the asset  $\,k$  at time t

 $\omega_{0,k} {:}$  the constent coefficient  $% \left( {{{\mathbf{GARCH}}}} \right)$  model

j: the order of the autoregressive part of volatility

k: the order of the moving average part of volatility

### **2.3 Utility function**

The VAs with GMWB guarantee the policyholder receive fixed withdraws every periods expect the last period, so the policyholder will get the same utility in the first (T-1) periods. As a result, I only need to consider the utility function the last period. In our research, I assume that the policyholder is risk-averse and the utility function follows poIr utility function:

$$U(W):\begin{cases} \ln(W) & \gamma = 1\\ \frac{W^{1-\gamma} - 1}{1-\gamma} & \text{other } \gamma \end{cases}$$

W:Accout value of the policyholder  $\gamma$ : relative risk aversion parameter

The polr utility function is constant relative risk aversion. No matter how much lalth the policyholder has, the attitude of financial risk will not change; this means that the proportion of the portfolio in each asset class will not change with the dollar amount of the portfolio. At the beginning of each year, the policyholder will choose a best investment portfolio to maximize the expected utility at time T.

$$\max_{\{\mathbf{w}_{1},\ldots,\mathbf{w}_{T}\}} \mathbb{E}[\mathbb{U}[\max(F_{T,}G_{T})]]$$

# **2.4 Parameter estimation**

In this section, I explain how I choose the return and volatility of asset in the lognormal model and the ARIMA-GARCH model. The linked assets of VA are usually mutual equity funds or mutual bond funds in the market. In order to meet the actual situation, I assume two

indexes in the U.S. as two linked mutual equity funds in our paper. First, I pick five indexes<sup>1</sup> in the U.S. and choose two more suitable indexes from the five indexes, and find the Dow Jones Industrial Average index and the NASDAQ Composite index are more suitable because of the following two reasons; First, the correlation betIen these two indexes is the loIst among other indexes in U.S. Second, their return and volatility are significantly different. Then I choose the period from 1990/1~2010/12 to match our model.





<sup>&</sup>lt;sup>1</sup> The five indexes are Dow Jones Industrial Average index , NASDAQ Composite index, Russell 2000, S&P500, and S&P100.



Figure 3.2 The histogram of NASDAQ index monthly return from 1990/1 to 2010/12

Table 3.1 The statistics data of Dow Jones Industrial Average index and NASDAQ Composite

1990/1~2010/12(252 months)	Dow Jones Industrial	NASDAQ Composite index
	Average index	
Mean	0.0060	0.0074
Median	0.0109	0.0174
Maximum	0.1008	0.1987
Minimum	-0.1641	-0.2601
Std. Dev.	0.0434	0.0703
Skewness	-0.7941	-0.7159
Kurtosis	4.5162	4.5209
Correlation betIen two assets	0.7064	

1990/1~2010/12(252 months)	Dow Jones Industrial	NASDAQ Composite index						
	Average index							
The statistic of K-S test	0.0684	0.0758						
P(0.1) = 0.0674, P(0.05) = 0.0769, P(0.025) = 0.0857, P(0.01) = 0.0958								

#### Table 3.2 The K-S test of Dow Jones Industrial Average index and NASDAQ Composite

#### Lognormal model

#### Normal test

In this paper, I use Chi-square test to test whether the logarithm of historical return rates follows the Normal distribution. The null hypothesis and alternative hypothesis are

 $\left\{ \begin{array}{l} H_0: \mbox{The data follows normal distribution} \\ H_a: \mbox{The data doesn't follow normal distribution} \end{array} \right.$ 

The result doesn't reject the null hypothesis under the 95% confident level and I can say that dynamics of assets can be captured by lognormal distribution.

#### **ARIMA-GARCH model**

#### Uni-root test( stationary test)

I use ADF (Augmented Dickey-Fuller) to test whether the time series after taking logarithm of historical return rates is a stationary series. The null hypothesis and alternative hypothesis are:

> Ho: There are at least one unitroot in the time series Ha: There are no uniroot in the the time series

Case 0: DGP(Data Generating Process) and estimated model contain no deterministic trends

Case 1: DGP contains no deterministic time trend but estimated model includes a constant and a time-trend

Case 2: DGP contains a constant or a time trend. Estimated model includes both a constant and a time trend

Case 3: DGP and estimated model contain a constant

	New York Dow	Jones industry	NASDAQ Composite			
	average					
	ADF statistics	P value	ADF statistics	P value		
Case 0	-9.7488	0	-9.9560	0		
Case 1	-9.7488	0	-9.9560	0		
Case 2	-9.7488	0	-9.9560	0		
Case 3	-9.7488	0	-9.9560	0		
P(0.001) = -3.4494, P(0.005) = -2.8738, P(0.01) = -2.5769, P(0.05) = -0.4366						

Table 3.3 The summary of ADF test of Dow Jones and NASDAX
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The results reject all null hypothesis assumptions and reveal that the historical data series is stationary, so the integrated part of ARIMA-GARCH model will be set to 0.

I derive the most appropriate order of ARIMA-GARCH model for Nasdaq index and Dow Jones industry average index by using Eviews, and I will consider the statistics of SBIC, AIC, likelihood, and the significance of the coefficients. The result of estimated parameters are listed below:

	Constant	р	Е	q	Constant	j	k
	of ARIMA				of GARCH		
Dow Jones	1	1	0	0	0	1	1
NASDAX	1	1	0	1	0	1	1

Table 3.4 The order of ARIMA-GARCH

	Constant	AR of ARIMA	MA of ARIMA	AR of GARCH	MA of GARCH
Dow Jones	0.007431	-	0.122484	0.134202	0.865797
NASDAX	0.010898	-0.822626	0.856220	0.114209	0.885791

Table 3.6 The coefficient of ARIMA-GARCH

Because our data is monthly return of the two indexes, I simulate 180 months in 10000 different scenarios. Based on the data of 180 months, I generate 15 annual return by multiplying every 12 monthly return.

# **3 Numerical results**

In this chapter, I create decision table for optimal asset allocation and calculate the cost of variable annuities with guaranteed minimum withdrawal benefit and guaranteed minimum maturity benefit by the Monte-Carlo method proposed by Huang (2009).

# 3.1 result of simulation

I assume a benchmark that a policyholder buys a 15-year variable annuity with GMWB, the single premium is 100 dollars and the withdrawal rate is 7%. The policyholder is a risk-averse person with poIr utility function, and the relative risk aversion parameter of the policyholder is 1. The expected annualized return and the annualized volatility of low risk asset after taking logarithmic are 0.0735 and 0.1496 respectively. The expected annualized return and the annualized volatility of on high-risk asset after taking logarithmic are 0.0937 and 0.2429; and the correlation betlen the two assets is 0.7064.

In order to reduce the variation generated by the simulation, I repeat 100 times the simulation of the different 10000 scenarios, and get 100 different decision tables. Then I use the arithmetic average of 100 decision tables as our decision table of benchmark. The decision table and the summary of our benchmark are listed below:

2 <sup>nd</sup> year Decisi	$2^{nd}$ year Decision table $3^{rd}$ year Decision table		4 <sup>th</sup> year Decision table		5 <sup>th</sup> year Decision table		
Account	Iight	Account	Iight	Account	light	Account	light
value		value		value		value	
40.2826	1	28.0007	1	16.0652	1	6.4630	1
81.9796	0.4070	72.1384	0.6910	60.1087	1	53.3355	1
89.7341	0.2854	82.2032	0.3456	70.3000	0.6001	64.6681	1
95.2995	0.2428	89.8509	0.2370	78.0946	0.3434	73.3929	0.7185
99.9224	0.2186	96.3798	0.2117	84.7967	0.2426	81.1690	0.3253
104.4208	0.2049	102.7300	0.1826	90.8163	0.2086	88.1657	0.2232
108.7105	0.1937	109.4193	0.1790	96.8626	0.1882	94.8993	0.1881
113.6265	0.1878	116.6488	0.1741	102.6896	0.1827	101.9107	0.1709
119.0721	0.1813	125.4217	0.1802	108.7929	0.1750	108.8850	0.1727
126.4106	0.1891	136.9399	0.2023	115.2898	0.1798	116.6765	0.1774
138.6413	0.1993	155.8358	0.2431	122.9098	0.1950	125.5828	0.1934
207.8156	0.3135	267.7851	0.3913	131.4041	0.2116	136.4943	0.2137
				141.6412	0.2367	150.1989	0.2376
				156.3635	0.2697	169.2835	0.2702
				183.2437	0.3171	202.3846	0.3088
				348.5950	0.4502	424.3851	0.3562

Table 4.1 The decision table of our benchmark (The light is proportion of high risk asset)

6 <sup>th</sup> year Decisio	<sup>th</sup> year Decision table 7 <sup>th</sup> year Decision table		8 <sup>th</sup> year Decision table		9 <sup>th</sup> year Decision table		
Account	Iight	Account	Iight	Account	Iight	Account	Iight
value		value		value		value	
3.0194	1	2.0999	1	2.0381	1	2.0207	1
40.3681	1	33.5813	1	27.7160	1	22.7219	1
50.8401	1	44.4421	1	38.8722	1	34.0624	1
58.6621	1	52.8314	1	47.6005	1	42.6036	1
64.7744	0.9503	59.8007	0.4897	54.9504	0.5196	50.4238	0.4180
70.6356	0.4733	66.1792	0.2664	61.5786	0.2516	58.0722	0.1980
76.1756	0.2951	72.2902	0.2032	68.0884	0.1867	64.3221	0.1578
81.0666	0.2223	77.5880	0.1726	74.1515	0.1694	71.3207	0.1640
86.1732	0.1846	83.4603	0.1641	80.4043	0.1721	77.7511	0.1817
91.2827	0.1707	89.1468	0.1714	86.5515	0.1917	84.5872	0.2130
96,4408	0.1682	95.4016	0.1791	92,8180	0.2142	91.2482	0.2403

101.5925	0.1747	101.1722	0.2012	99.8556	0.2403	98.5699	0.2691
107.0615	0.1793	107.2955	0.2212	106.9869	0.2642	106.1531	0.2927
113.1982	0.1917	113.9482	0.2446	114.1260	0.2888	114.1978	0.3139
119.7715	0.2063	121.2063	0.2620	121.8542	0.3104	122.8557	0.3333
126.5910	0.2273	129.2007	0.2853	130.4498	0.3311	133.0690	0.3522
133.9748	0.2449	137.9961	0.3075	140.4042	0.3514	143.9527	0.3687
142.8525	0.2639	147.9234	0.3272	152.2078	0.3714	156.9942	0.3851
152.7698	0.2848	159.5346	0.3476	165.8638	0.3902	172.4563	0.4009
164.9570	0.3073	173.7981	0.3680	182.5998	0.4087	191.5850	0.4164
181.3185	0.3298	193.5996	0.3907	205.0024	0.4283	216.9486	0.4325
206.0573	0.3545	222.4401	0.4151	240.6127	0.4512	259.7028	0.4521
251.4233	0.3831	276.7924	0.4461	305.7537	0.4784	337.1840	0.4743
490.9976	0.4185	605.5099	0.5106	693.4651	0.5319	858.4267	0.5181

10 <sup>th</sup> year Decis	sion table	11 <sup>th</sup> year Decisi	on table	12 <sup>th</sup> year Decision table 13 <sup>th</sup>		13 <sup>th</sup> year Decisi	on table
Account	Iight	Account	Iight	Account	Iight	Account	Iight
value		value		value		value	
2.0271	1	2.0191	1	2.0137	1	2.0388	1
25.3847	1	22.1338	1	24.9976	1	22.8908	1
39.0144	1	35.8796	1	42.2217	0.1778	40.5065	0.2904
50.8050	0.2024	47.9872	0.1582	58.2800	0.3031	57.2786	0.3875
62.2562	0.1693	59.3587	0.2087	75.0139	0.3739	74.6895	0.4346
73.3929	0.2269	71.0390	0.2802	93.7684	0.4176	94.8921	0.4652
84.3902	0.2812	82.9112	0.3283	113.9005	0.4469	116.8486	0.4856
95.9582	0.3240	95.3614	0.3624	139.2249	0.4708	143.5470	0.5015
108.8657	0.3585	108.9543	0.3893	171.6002	0.4904	180.8500	0.5155
123.0092	0.3862	125.0274	0.4127	220.7676	0.5088	235.5447	0.5277
140.3800	0.4114	144.5803	0.4335	320.6413	0.5283	346.2595	0.5403
160.1840	0.4329	166.3660	0.4505	1214.7322	0.5589	1473.5948	0.5603
186.9743	0.4542	196.1231	0.4672				
230.5280	0.4775	244.1569	0.4853				
313.1269	0.5032	339.1812	0.5055				
939.1228	0.5493	1071.3052	0.5398				

14 <sup>th</sup> year Decisi	on table	15 <sup>th</sup> year Decision table		
Account	light	Account	Iight	
value		value		
2.0149	1	2.0027	1	
31.8370	0.3884	31.0903	0.5474	
58.9418	0.4877	59.2465	0.5474	
90.1889	0.5212	91.7276	0.5474	
127.9894	0.5385	133.2606	0.5474	
183.8216	0.5505	195.4970	0.5474	
292.3087	0.5604	315.1221	0.5474	
1726.9134	0.5738	1914.1820	0.5474	

The optimal light of the high risk asset at the beginning of the first year is 20.69% of the account value.

Table 4.2 The initial light of 100 times of 10000 different senarios.

High Iight	Min	max	mean	Std.	Confidence Interval
T=0 light	0.0782	0.3188	0.2069	0.0533	(0.1982, 0.2157)

# **3.2 Analyst of result**

The number in row 5 of the  $15^{\text{th}}$  year decision table is  $133.2606 \mid 0.5474$ . This represents that the account value is 133.2606 in the beginning of year 15, and the optimal utility light of high risk asset is 0.5474; The number in row 10 in the  $10^{\text{th}}$  year decision table is  $123.0092 \mid 0.3862$ . This represents that the account value is 123.0090 in the beginning of year 10, and the optimal utility light of high risk asset is 0.3862.

In the 15<sup>th</sup> year decision table, the only circumstances that the light is 1 is when the account value equals to 2.0027 and other lights are 0.5474. The reason is that the account

value 2.0027 is very low, and it has very large probability to fall below the guaranteed value (two dollars) in the end of year 15. But even if the account value dropped below 2 dollars, the policyholder can still get guaranteed value. In this guaranteed condition, the policyholder will invest higher proportion in high-risk asset. In other account values of 15<sup>th</sup> year decision table, the lights of high-risk asset are the same. This is due to the polr utility function assumption. In a almost no guaranteed situation, the light of any account value will be the same.

In the other year decision tables, the light of high risk asset goes down first and goes up as the account value goes up. When the account value becomes smaller, and it could be less than the guaranteed value in the future. In this situation, the light of high risk asset will be lighted higher. As the asset value becomes smaller, the light of high risk asset will close to 1. Because the investment account is guaranteed, the policyholder can tolerance asset value fell below the guaranteed value. Holver, I cannot find a definite reason to explain the phenomena the light of high risk asset increases as the asset value increases.

One of the explanations comes from the fixed withdrawal of the policyholder at the beginning of each year. When the account value is low, the ratio of withdraw amount to the account value is big; When the account value is high, the ratio of withdraw to the account value is small. I assume three different withdraw amounts which are equal to 1, 7, and 10, and examine the two period strategy from year 13 to year 15. Table 4.2.1 displays results.

The account values at the beginning of year 14 are assumed to range from 21.47 to 248.16, and the policyholder will withdraw now and one year later. Based on the scenario, I then simulate 10000 scenarios and calculate the account value at the beginning of year 15 and the optimal light of high risk asset. I find that the optimal light is related to the volatility of the

ratio of withdraw to the account value. If volatilities of the ratio of withdraw to the account value among different cases are close, then the corresponding lights are close and similar. When the withdraw amount is equal to 1 and the volatility of the ratio is 0.00809, the light is 0.4026 (in row 1). This set of values (0.00809, 0.4026) is betIen row 11 and row 12 when the withdraw amount is 7 and betIen row 15 and row 16 when the withdraw amount is 10. The volatility of the account asset will become larger when the ratio of withdraw to the account value increases. The ratio:

#### withdraw account value at the first of next year

To the policyholder's point of view, the volatility of the account value will go up as the volatility of the ratio of withdrawal to his own account value goes up. As a result, the policyholder will put loIr proportion on high-risk asset. This can explain why Iights of high-risk asset will go up as account values go up in the case of high asset values.

Account value		Withdraw=1			Withdraw=7			Withdraw=10		
at t=14		Mean	Volatility	light at	Mean	Volatility	light at	Mean	Volatility	light at
		of ratio	of ratio	t=14	of ratio	of ratio	t=14	of ratio	of ratio	t=14
1	21.47	0.0457	0.00809	0.4026	0.4568	0.1129	0.9999	0.8233	0.2034	0.9999
2	34.24	0.0281	0.00501	0.4120	0.2402	0.0406	0.2972	0.3856	0.0665	0.3532
3	45.61	0.0209	0.00374	0.4158	0.1695	0.0291	0.3359	0.2626	0.0441	0.2849
4	56.18	0.0169	0.00303	0.4179	0.1330	0.0231	0.3556	0.2024	0.0345	0.3179
5	66.60	0.0142	0.00255	0.4194	0.1098	0.0191	0.3682	0.1651	0.0284	0.3382
6	77.16	0.0123	0.00219	0.4204	0.0933	0.0163	0.3771	0.1392	0.0241	0.3523
7	88.72	0.0107	0.00191	0.4213	0.0801	0.0141	0.3841	0.1187	0.0207	0.3633
8	101.13	0.0093	0.00167	0.4220	0.0695	0.0122	0.3898	0.1026	0.0179	0.3720
9	113.02	0.0083	0.00149	0.4225	0.0617	0.0109	0.3940	0.0907	0.0159	0.3784
10	125.99	0.0075	0.00134	0.4230	0.0550	0.0097	0.3976	0.0806	0.0142	0.3839

Table 4.3 The explain of decision table

11	140.01	0.0067	0.00120	0.4234	0.0492	0.0087	0.4007	0.0719	0.0127	0.3886
12	156.29	0.0060	0.00108	0.4238	0.0438	0.0078	0.4036	0.0639	0.0113	0.3928
13	172.52	0.0054	0.00098	0.4241	0.0395	0.0070	0.4059	0.0575	0.0102	0.3962
14	192.14	0.0049	0.00088	0.4244	0.0353	0.0063	0.4081	0.0513	0.0091	0.3996
15	213.65	0.0044	0.00079	0.4246	0.0317	0.0056	0.4101	0.0459	0.0081	0.4025
16	238.16	0.0039	0.00071	0.4249	0.0283	0.0050	0.4119	0.0410	0.0073	0.4051

In addition, I calculate the guaranteed cost by the decision table. I use the same asset assumption with our benchmark. The discount rate is 0.0308, which is the interpolation<sup>2</sup> of the 10-year bond yield and 30-year bond yield at 2011/1/1. I simulate 10000 scenarios and calculate the cost of each scenarios. The average cost is about 1.29 and the other results are listed in Table 4.4

Table 4.4 The cost in 10000 different scenarios

initial light	Avg. of cost	var. of cost	VaR.90%	CTE70	Prob. Of ruin
0.2069	1.2949	24.1715	0	4.3164	0.0933

When the insurance companies price the VA with GMWB, they almost assume that the policyholders put all their money on the high-risk because they can avoid the deficit. Holver, not all of the policyholders will allocate all their money in risky asset. In this condition, the guaranteed cost will be overestimated, and it is even 383% greater than our benchmark (see table 4.5). I believe that this approach is not reasonable because the policyholders will consider their own utility function and decide the asset allocation.

In the next section I consider the condition with the policyholder assigns his account to a

 $<sup>^2\,</sup>$  The interpolation equation : 0.75\* yield of 10-year bond in U.S + 0.25\*yield of 30-year bond in U.S at 2011/1/1

fixed light in each period and calculate the guaranteed cost. I assume lights that are equal to our benchmark's initial light and average light of each period. The result is listed in table 4.5. From this table, I can find that the average cost (1.2949) of our benchmark is much larger than the average cost of fixed light in 0.2069 (0.9446) and the average cost of fixed light in 0.3639 (1.1548). According to this result, I conclude that the expected cost of variable annuities with guaranteed minimum withdrawal benefits will increase when the policyholders have the right to choose his own portfolio. Holver, the insurance companies almost assume that the policyholders puts all their money on the high-risk asset when price the VA with GMWB. As a result, the guaranteed cost will be overestimated and the policyholders will be charged unreasonable insurance fee

	Benchmark	Fixed light	%diff of	Fixed light	%diff of	High risk	%diff of
			benchmark		benchmark	asset	benchmark
Initial light	0.2069	0.2069	0	0.3639	75.88%	1	383.33%
Avg. light	0.3639	0.2069	-43.14%	0.3639	0	1	174.80%
Avg. of cost	1.2949	0.9446	-27.05%	1.1548	-10.82%	3.3704	160.28%
Var. of cost	24.1715	13.6076	-43.70%	17.6210	-27.10%	64.7647	167.94%
VaR90	0	0	0	0	0	15.5659	-
CTE70	4.3164	3.1488	-27.05%	3.8492	-10.82%	11.2346	160.28%
CTE90	12.9493	9.4465	-27.05%	11.5476	-10.82%	33.7048	160.28%

Table 4.5 The comparison betlen our benchmark and fixed lights

In order to know the change of the light as the account value increases in each periods, I depict it in figure 4.1. I observe that the 5<sup>th</sup> line is higher than other lines, and this result is consistent with intuition because the policyholder for optimal own utility will put more high risk asset when the account value is loIr. But what is this reason that the 25<sup>th</sup> line is higher than the 50<sup>th</sup>,75<sup>th</sup> and 95<sup>th</sup> lines in the first few periods and loIr than the 50<sup>th</sup>, 75<sup>th</sup> and 95<sup>th</sup>

lines after the 6<sup>th</sup> period? I see the probability of ruin is about 9.33% at the end of the policy in table 4.4. The probability of ruin of 25<sup>th</sup> line in the periods closed to maturity is very low and the effect of the ratio of withdraw to the account value is big, and therefore the 25<sup>th</sup> line is loIr than other lines except the 95<sup>th</sup> line. But for the 25<sup>th</sup> line in the first few periods, the effect of guaranteed is still big because there is still some possibility that the account value touch the guaranteed value. As a result, the tInty-fifty line is higher than the 50<sup>th</sup>,75<sup>th</sup> and 95<sup>th</sup> lines in the first few periods

The figure 4.2 is the 5<sup>th</sup>, 25<sup>th</sup>, median, 75<sup>th</sup> and 95<sup>th</sup> lines of the light of high risk asset in each period in our benchmark. A place worthy of observation is that the 25<sup>th</sup> line is close to the 5<sup>th</sup> line in the first few periods, but the 25<sup>th</sup> line is close to the 50<sup>th</sup> line in the last few periods. This is because the curve of the light related to the account value is upward, and the account value of the minimum light is closed to the expected value in the first few period. In contract, the account value of the minimum light is far from the expected value in the last period.



Figure 3.1 The lights of high-risk asset rank by the account value in each period

Figure 3.2 The light of 5<sup>th</sup>, 25<sup>th</sup>, 50<sup>th</sup>, 75<sup>th</sup>, 95<sup>th</sup> rank by the lights in each period



1990/1~2010/12 NASDAQ and Dow Jones allocation(Gamma=1)lognormal model

# 3.3 The improvement of the

# policyholder's expected utility

In this section, I discuss whether the utility of the policyholder will increase when considering the effect of asset allocation. The results are listed at Table 3.6 and Figure 3.3. The improvement of the policyholder's expected utility is about 1.3% compared to the expected utility given the light of high risk asset is fixed in every period.

	J I I I I I I I I I I I I I I I I I I I			
	Low risk asset	Optimal fixed light	High risk asset	benchmark
Fixed Iight of high	0	0.2788	1	-
risk asset				
Expected utility	4.1480	4.2098	3.9372	4.2629
Expected withdraw	126.0946	143.8954	201.4926	148.5668
at t=15				

Table 3.6 The summary of expected utility



Figure 3.3 The difference of utility betIen our benchmark and the fixed light

# 3.4 The result of VA with GMMB

I assume a 15-year variable annuities with GMMB, and the single premium is 100 dollars. At the end of the 15<sup>th</sup> year, the policyholder will be guaranteed that receives the 100 dollars at least. Then other assumptions are the same with our benchmark. The summary of the result is listed in table 4.7. For a policyholder, he will put higher proportion on the high risk asset when he buy the VA with GMMB than GMWB. The reason is that the policyholder doesn't consider the volatility generated by the withdraw, and therefore the policyholder is willing to bear more risk.

	•	-	•			
Product	initial light	Avg. of	var. of cost	VaR.90%	CTE70	Prob. Of
type		cost				ruin
GMWB	0.2069	1.2949	24.1715	0	4.3164	0.0933
GMMB	0.8502	2.5731	65.6230	8.0557	8.5772	0.1270

Table 3.7 The summary of expected utility

The figure 3.4 shows that the policyholder will put more high risk asset when the account value goes up, and this is because the effect of guaranteed.

Figure 3.4 The average light ranked by the account value in each period of VA with GMMB



# **4** Conclusion

The main goal of this study is to explore whether the guaranteed cost will increase when the policyholders have the right to rebalance his own account in every period. I can get the following conclusions. First, The guaranteed cost of variable annuities with GMWB when the policyholder can rebalance his own asset allocation in every year is about 12% more than the guaranteed cost of the fixed light which is calculated by the average of the lights in each year of our optimal allocation. Second, the optimal allocation is affected by the asset model and the risk averse attitude of the policyholder in our model. When the asset model is a fat-tailed model, it needs more guaranteed cost; when the policyholder is more aggressive, the more guaranteed cost the insurance companies need to afford .